

Constructing functional linear filters

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Outline

- 1 Measurable linear transformations
- 2 Innovation of ARMAH processes
- 3 Inverse problems
- 4 Computing linear filters in Hilbert spaces
- 5 Statistics...

Linear prediction in large dimensions

Example: evolution of US Economy based on simultaneous observation of 500 series

Goal: Explicit expression of the Best Linear Predictor in a function space

Difficulty: The associated linear operator is, in general, NOT continuous

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Hilbert spaces

H : real separable Hilbert space with norm $\|\cdot\|$ and scalar product $\langle \cdot, \cdot \rangle$

\mathcal{L} : space of continuous linear operators from H to H with its usual norm $\|\cdot\|_{\mathcal{L}}$

$L_H^2 = L_H^2(\Omega, \mathcal{A}, P)$: Hilbert space of (classes of) random variables defined on the probability space (Ω, \mathcal{A}, P) and with values in (H, \mathcal{B}_H) , scalar product

$$[X, Y] = E \langle X, Y \rangle ; X, Y \in L_H^2.$$

In the following all the random variables are supposed to be **centered**.

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Linearly closed subspaces

A linear subspace \mathcal{G} of L_H^2 is said to be **linearly closed (LCS)** if \mathcal{G} is closed in L_H^2 and $X \in \mathcal{G}, I \in \mathcal{L}$ implies $I(X) \in \mathcal{G}$.

X and Y in L_H^2 are said to be *weakly orthogonal* ($X \perp Y$) if $E \langle X, Y \rangle = 0$ and *strongly orthogonal* if $C_{X,Y} = 0$ where

$$C_{X,Y}(x) = E(\langle X, x \rangle Y), \quad x \in H$$

is the *cross-covariance operator* of X and Y .

Y weakly orthogonal to \mathcal{G} implies Y strongly orthogonal to \mathcal{G} .

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Measurable linear transformation

Let μ be a Probability on (H, \mathcal{B}_H) . An application λ is said to be a **μ -measurable linear transformation (μ -MLT)** if λ is measurable and linear on a linear space S such that $\mu(S) = 1$.

It is equivalent to say that there exists a sequence $(I_k, k \geq 1)$ in \mathcal{L} such that

$$I_k(x) \xrightarrow[k \rightarrow \infty]{} \lambda(x), \quad x \in S.$$

(cf Mandelbaum (1984)).

λ is, in general, NOT continuous, example:

$$\lambda(x) = x'$$

In the following λ always denotes a MLT.

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The gaussian case

In the gaussian case one has a more precise property:

Lemma

Let X be a H -valued gaussian random variable and let \mathcal{G}_X be the LCS generated by X . If λ is P_X -MLT there exists $(I_k, k \geq 1)$ in \mathcal{L} such that

$$E \|I_k(X) - \lambda(X)\|^2 \xrightarrow[k \rightarrow \infty]{} 0,$$

it follows that $\lambda(X) \in \mathcal{G}_X$.

An example

$H = L^2(\mathbb{R})$, $(h_j, j \geq 0)$ the orthonormal basis of Hermite functions, set

$$X = \sum_{j=0}^{\infty} \xi_j h_j$$

where the ξ_j 's are real independent and such that

$$P(\xi_j = -a_j) = P(\xi_j = a_j) = p_j, \quad j \geq 1$$

with $p_j < \frac{1}{2}$, $\sum_j p_j < \infty$ and $a_j > 0$, $\sum_j p_j a_j^2 < \infty$. Then $P(X \in S) = 1$

where S is the **linear space of polynomials** with weight $\exp(-\frac{t^2}{2})$, $t \in \mathbb{R}$ and if $\lambda(x) = x'$ and $l_k(x)(t) = \frac{x(t+1/k) - x(t)}{1/k}$, $t \in \mathbb{R}$, $k \geq 1$, then

$$2k \|l_k(x) - \lambda(x)\| \xrightarrow[k \rightarrow \infty]{} \|\lambda^2(x)\|, \quad x \in S.$$

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Projection on a LCS

The link between MLT and LCS appears in the following statement

Proposition

Let \mathcal{G}_X be the LCS generated by X and Π^X its orthogonal projector in L^2_H . Then, for each Y in L^2_H , there exists a P_X -MLT λ_0 such that

$$\Pi^X(Y) = \lambda_0(X).$$

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Continuity

The next proposition underscores a special case where λ_0 is **continuous**:

Proposition

The following statements are equivalent

- a) *There exists $\alpha \geq 0$ such that $\|C_{X,Y}(x)\| \leq \alpha \|C_X(x)\|$, $x \in H$,*
- b) *There exists $l_0 \in \mathcal{L}$ such that $C_{X,Y} = l_0 C_X$,*
- c) *$\Pi^X(Y) = l_0(X)$.*

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- c) $\Pi^X(Y) = l_0(X)$.

Innovation

A H - **white noise** is a sequence $(\varepsilon_n, n \in \mathbb{Z})$ of strongly orthogonal H -valued random variables such that $E \|\varepsilon_n\|^2 = \sigma^2 > 0$ and $E\varepsilon_n = 0, n \in \mathbb{Z}$.

A **weakly stationary process** *in* H satisfies

$$C_{X_{n+h}, X_{m+h}} = C_{X_n, X_m}, \quad n, m, h \in \mathbb{Z}.$$

$(\varepsilon_n, n \in \mathbb{Z})$ is the **innovation** of $(X_n, n \in \mathbb{Z})$ if

$$X_{n+1}^* = X_n + \varepsilon_{n+1}, \quad n \in \mathbb{Z},$$

where X_{n+1}^* is the best linear predictor of X_{n+1} given X_n, X_{n-1}, \dots

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ARH(1)

Set \mathcal{M}_n be the LCS generated by X_n, X_{n-1}, \dots . A stationary process is an **autoregressive process of order 1** in H (ARH(1)) if

$$\Pi^{\mathcal{M}_{n-1}}(X_n) = \Pi^{\mathcal{G}_{X_{n-1}}}(X_n)$$

Hence

$$X_n = \lambda_n(X_{n-1}) + \varepsilon_n, \quad n \in \mathbb{Z},$$

where λ_n is MLT and (ε_n) is the innovation.

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Innovation of an ARH(1)

Proposition

Suppose that the equation

$$X_n = \lambda(X_{n-1}) + \varepsilon_n, \quad n \in \mathbb{Z} \quad (1)$$

has a solution such that $\lambda : S \mapsto S$ is P_{X_n} -MLT for all n , $\lambda^j(X_{n-j}) \in \mathcal{G}_{X_{n-j}}$ and $\lambda^j(\varepsilon_{n-j}) \in \mathcal{G}_{\varepsilon_{n-j}}$, $j \geq 1$, then if

$$\frac{1}{k} \sum_{j=1}^k E \|\lambda^j(X_{n-j})\|^2 \xrightarrow[k \rightarrow \infty]{} 0, \quad n \in \mathbb{Z}$$

(1) has a unique stationary solution given by

$$X_n = \lim_{k \rightarrow \infty} (L_H^2) \sum_{j=0}^{k-1} \left(1 - \frac{j}{k}\right) \lambda^j(\varepsilon_{n-j}), \quad n \in \mathbb{Z}$$

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and (ε_n) is the **innovation** of (X_n) .

Proof

The proof is based on the relation

$$X_n = \sum_{j=0}^{k-1} \left(1 - \frac{j}{k}\right) \lambda^j (\varepsilon_{n-j}) + \frac{1}{k} \sum_{j=1}^k \lambda^j (X_{n-j}).$$

The above condition is strictly weaker than the classical conditions like:

“ λ is continuous and there exists an integer j_0 such that $\|\lambda^j\|_{\mathcal{L}} < 1, j \geq j_0$.” (cf Bosq-Blanke 2007)

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Innovation of a MAH(1)

Proposition

Suppose that (X_n) is defined by

$$X_n = \varepsilon_n - \lambda(\varepsilon_{n-1}), \quad n \in \mathbb{Z}.$$

where $\lambda : H_1 \mapsto H_1$ is P_{ε_n} -MLT for all n , with $\lambda^j(X_{n-j}) \in \mathcal{G}_{X_{n-j}}$, $\lambda^j(\varepsilon_{n-j}) \in \mathcal{G}_{\varepsilon_{n-j}}$, $j \geq 1$, $n \in \mathbb{Z}$, then, if

$$\frac{1}{k^2} \sum_{j=1}^k E \|\lambda^j(\varepsilon_{n-j})\|^2 \xrightarrow[k \rightarrow \infty]{} 0,$$

(ε_n) is the *innovation* of (X_n) and

$$\varepsilon_n = \lim_{k \rightarrow \infty} (L_H^2) \sum_{j=0}^{k-1} \left(1 - \frac{j}{k}\right) \lambda^j(X_{n-j}).$$

“Roots of modulus 1”

The condition is weak. In particular if λ is continuous and such that

$$\|\lambda^j\|_{\mathcal{L}} \leq 1, j \geq 1$$

the above Proposition holds. A simple example is

$$X_n = \varepsilon_n - \Pi^G(\varepsilon_{n-1}), n \in \mathbb{Z}$$

where G is a closed subspace of H and Π^G its orthogonal projector.

If the MA is real, it corresponds to **roots of modulus 1**.

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Example

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In $L^2[0, 1]$ consider the white noise

$$\varepsilon_n(t) = \sum_{i=0}^{\infty} \xi_{ni} \frac{t^i}{i!}, \quad t \in [0, 1], \quad n \in \mathbb{Z}$$

where (ξ_{ni}) is a sequence of real independent random variables such that, for all n , $\xi_{ni} \sim \mathcal{N}(0, \sigma_i^2)$ where $0 < \sum_{i=1}^{\infty} \sigma_i^2 < \infty$. Set

$$X_n(t) = \varepsilon_n(t) - \varepsilon'_{n-1}(t)$$

then (ε_n) is the innovation.

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The mixed case

Proposition

Consider the ARMAH (1,1) process defined as

$$\varepsilon_n - l(\varepsilon_{n-1}) = X_n - \rho(X_{n-1}), \quad n \in \mathbb{Z}$$

where (ε_n) is a H -white noise and l and ρ belong to \mathcal{L} ; suppose that

$$\frac{1}{k^2} \sum_{j=0}^k \|\mu^j\|_{\mathcal{L}}^2 \xrightarrow[k \rightarrow \infty]{} 0 \quad \text{and that} \quad \frac{1}{k} \sum_{j=0}^k \|\rho^j\|_{\mathcal{L}}^2 \xrightarrow[k \rightarrow \infty]{} 0$$

then, if that equation has a stationary solution, it is given by

$$X_n = \lim_{k \rightarrow \infty} (L_H^2) \sum_{j=0}^{k-1} \left(1 - \frac{j}{k}\right) \rho^j(\varepsilon_{n-j} - l(\varepsilon_{n-j-1})), \quad n \in \mathbb{Z}$$

and (ε_n) is the innovation of (X_n) .

Compound Ornstein-Uhlenbeck process

Example

Consider the Hilbert space $H = L^2([0, 1], \mathcal{B}_{[0,1]}, \mu)$ where μ is the sum of Lebesgue measure and Dirac measure at the point 1. Set

$$\varepsilon_n(t) = \int_n^{n+t} \exp(-\theta(n+t-s)) dW(s), \quad t \in [0, 1], \quad n \in \mathbb{Z}, \quad (\theta > 0),$$

where W is a bilateral standard Wiener process. Put $l(x)(t) = x(t)$, and $\rho(x)(t) = \exp(-\theta t) \cdot x(1)$ $t \in [0, 1]$, $x \in H$. Then the process

$$\begin{aligned} X_n(t) = & \exp(-(\theta(n+t))) \int_{-\infty}^{n+t} \exp(\theta s) dW(s) \\ & - \exp(-\theta(n-1+t)) \int_{-\infty}^{n-1+t} \exp(\theta s) dW(s), \quad t \in [0, 1], \quad n \in \mathbb{Z} \end{aligned}$$

is a stationary ARMAH (1,1).

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is a stationary ARMAH (1,1).

Kalman-Bucy filter in H

Example

Consider the model

$$X_n = r(Y_n) + \varepsilon_n, \quad n \geq 1$$

$$Y_n = \rho(Y_{n-1}) + \eta_n, \quad n \geq 1$$

where (X_n) and (Y_n) are H -valued stationary processes and where (ε_n) and (η_n) are two strongly orthogonal white noises such that

$C_{\varepsilon_n, Y_n} = C_{\eta_n, Y_{n-1}} = 0$; ρ and r belong to \mathcal{L} . Then, if $r\rho = \rho r$, (X_n) is an ARMAH (1,1).

Other examples of Kalman-Bucy filter in H appear in Ruiz-Medina et al/ in a spatial framework.

Kalman-Bucy filter in H

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MAH(2)

Proposition

Consider a MAH(2) admitting the decomposition

$$X_n = \varepsilon_n - (\alpha + \beta)(\varepsilon_{n-1}) + \beta\alpha(\varepsilon_{n-2}), \quad n \in \mathbb{Z}$$

where (ε_n) is a white noise and $\alpha, \beta \in \mathcal{L}$ and suppose that

$$\frac{1}{k^2} \sum_{j=0}^k \|\alpha^j\|_{\mathcal{L}}^2 \xrightarrow[k \rightarrow \infty]{} 0$$

and

$$\frac{1}{k^2} \sum_{j=0}^k \|\beta^j\|_{\mathcal{L}}^2 \xrightarrow[k \rightarrow \infty]{} 0$$

then (ε_n) is the innovation of (X_n) .

Constructing the innovation

What about the case where the noise associated with the process is NOT the innovation?

The case of a MAH(1)

Proposition

Consider the MAH(1) given by

$$X_n = e_n - I(e_{n-1}), n \in \mathbb{Z}$$

where $I \in \mathcal{L}$ and (e_n) is a H -white noise. We suppose that I is symmetric, invertible, such that $\|(I^{-1})^{j_0}\|_{\mathcal{L}} < 1$ for some $j_0 \geq 1$. Moreover I and C_{e_0} commute.

Then, the innovation of (X_n) is defined as $\varepsilon_n = (I - I^{-1}B)^{-1}(I - IB)e_n$ where B is the backward operator ($B(x_n) = x_{n-1}$), convergence takes place in L_H^2 , and

$$X_n = \varepsilon_n - I^{-1}(\varepsilon_{n-1}), n \in \mathbb{Z}.$$

In addition one has

$$C_{\varepsilon_0} = I^2 C_{e_0}.$$

An example

Example

Suppose that

$$I = \sum_{i=1}^{\infty} a_i v_i \otimes v_i$$

where (v_i) is an orthonormal system in H and $1 < |a_1| \leq |a_2| \leq \dots \leq a < \infty$; and that

$$C_{e_0} = \sum_{i=1}^{\infty} c_i v_i \otimes v_i$$

then the above Proposition holds.

The case of an ARH(1)

Proposition

Consider the equation $X_n = r(X_{n-1}) + \eta_n$ $n \in \mathbb{Z}$
 where (η_n) is a H -white noise and $r \in \mathcal{L}$, and suppose that

$$\exists r^{-1} : \frac{1}{k} \sum_{j=1}^k \|r^{-j}\|_{\mathcal{L}}^2 \xrightarrow[k \rightarrow \infty]{} 0,$$

then it has a stationary solution given by

$$X_n = - \lim_{k \rightarrow \infty} (L_H^2) \sum_{j=1}^k \left(1 - \frac{j-1}{k}\right) r^{-j}(\eta_{n+j}), \quad n \in \mathbb{Z}.$$

If, in addition, $r^{-1}C_{X_0}$ is symmetric and $C_{X_0}(Is - (r^*)^{-2}) \neq 0$, then the innovation of (X_n) is

$$\varepsilon_n = X_n - r^{-1}(X_{n-1}) \quad n \in \mathbb{Z}$$

Starting from the best predictor

Principle: Given the best linear predictor (BLP) find the associated model.

Choice: **Extended exponential smoothing** in H :

$$X_{n+1}^* = \alpha \left(\sum_{j=0}^{\infty} \beta^j (X_{n-j}) \right),$$

where α and β belong to \mathcal{L} and $\alpha\beta = \beta\alpha$. Then one has

$$X_{n+1}^* = \alpha(X_n) + \beta(X_n^*).$$

Associated model

Proposition

Suppose that $\|\beta^{j_0}\|_{\mathcal{L}} < 1$ and $\|(\alpha + \beta)^{j_0}\|_{\mathcal{L}} < 1$ for some integer j_0 , and that $\alpha \neq 0$. If (X_n) is a regular zero-mean stationary process with innovation (ε_n) and such that the BLP is

$$X_{n+1}^* = \alpha \left(\sum_{j=0}^{\infty} \beta^j (X_{n-j}) \right)$$

where $\alpha\beta = \beta\alpha$, then (X_n) is an ARMAH (1,1):

$$X_n - (\alpha + \beta)(X_{n-1}) = \varepsilon_n - \beta(\varepsilon_{n-1}), \quad (2)$$

Conversely, if (X_n) satisfies 2, then X_{n+1}^* is BLP for every n .

(X, Y) in $G \times H$ real separable Hilbert spaces with spectral decompositions:

$$C_X = \sum_{i \in I} \alpha_i v_i \otimes v_i \quad (\alpha_i > 0, \sum_{i \in I} \alpha_i < \infty)$$

and

$$C_Y = \sum_{j \in J} \beta_j w_j \otimes w_j \quad (\beta_j > 0, \sum_{j \in J} \beta_j < \infty)$$

I and J are finite or infinite. Let $\mathcal{L}(G, H)$ be the space of continuous linear operators from G to H . Set

$$\mathcal{F}_X = \text{sp} \{l(X), l \in \mathcal{L}(G, H)\}$$

where the closure is taken in $L_H^2 = L_H^2(\Omega, \mathcal{A}, P)$.

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The best linear predictor

Proposition

The best linear predictor (BLP) of Y given X is

$$\lambda_0(X) = \sum_{i \in I, j \in J} \gamma_{ij} (v_i \otimes w_j)(X) \quad (L_H^2)$$

where

$$\gamma_{ij} = \frac{E(\langle X, v_i \rangle_G \langle Y, w_j \rangle_H)}{E \langle X, v_i \rangle_G^2}, \quad i \in I, j \in J$$

Proof

The proof uses the fact that

$$U_{ij} = \frac{\langle X, v_i \rangle_G}{\sqrt{\alpha_i}} \cdot w_j \quad i \in I, j \in J,$$

is an orthonormal system in L_H^2 .

Continuity

λ_0 is a P_X -MLT. Continuity of λ_0 appears in the next statement

Proposition

If there exists $l_0 \in \mathcal{L}(G, H)$ such that

$$C_{X,Y} = l_0 C_X,$$

then the best linear predictor takes the form

$$l_0(X) = \sum_{i \in I} \frac{\langle X, v_i \rangle_G}{\alpha_i} C_{X,Y}(v_i)$$

Converse

Proposition

If there exists $l_0 : G \mapsto H$ such that

$$l_0(x) = \sum_{i=1}^{\infty} \frac{\langle x, v_i \rangle_G}{\alpha_i} C_{X,Y}(v_i), \quad x \in H, \quad (H)$$

then $l_0 \in \mathcal{L}(G, H)$ and $C_{X,Y} = l_0 C_X$.

The gaussian case

In the gaussian case a similar result can be obtained without continuity assumption:

Proposition

If $G = H$ and the vector (X, Y) is gaussian then the conditional expectation $E(Y|X)$ and the BLP coincide and have the form

$$E(Y|X) = \lambda_0(X) = \sum_{i \in I} \frac{\langle X, v_i \rangle_G}{\alpha_i} C_{X, Y}(v_i)$$

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Proof

The proof uses the fact that the sequence

$$E(\langle Y, y \rangle_H | (\langle X, v_1 \rangle_G, \dots, \langle X, v_m \rangle_G)) = \sum_{i=1}^m \frac{E(\langle X, v_i \rangle_G \langle Y, y \rangle_H)}{E(\langle X, v_i \rangle_G^2)} \langle X, v_i \rangle_G,$$

$m \geq 1$, is a martingale in L_H^2 .

Basis of a LCS

The final statement is useful for computing a BLP

Proposition

The LCS \mathcal{G}_X of L_G^2 has the orthonormal basis

$$\mathcal{B} = \left\{ \frac{\langle X, v_i \rangle_G}{\alpha_i^{1/2}} v_j, i \in I, j \in I \right\} \cup \left\{ \frac{\langle X, v_i \rangle_G}{\alpha_i^{1/2}} u_j, i \in I, j \in J' \right\}$$

where

$$C_X = \sum_{i \in I} \alpha_i v_i \otimes v_i \quad (\alpha_i > 0, \sum_{i \in I} \alpha_i < \infty)$$

and $(u_j, j \in J')$ is an orthonormal basis of the orthogonal complement of the closed subspace of G generated by $(v_i, i \in I)$.

(cf Bosq-Mourid (2012)).

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Applications: model with noise

Consider the model

$$X = r(Y) + \varepsilon$$

with $r \in \mathcal{L}(H, G)$ and $C_{Y, \varepsilon} = 0$, where only X is observed, Then

$$C_{X, Y} = C_Y r^*$$

hence

$$\lambda_0(X) = \sum_{i,j} \frac{\beta_j}{\alpha_i} \langle v_i, r(w_j) \rangle_H \langle X, v_i \rangle_G w_j.$$

Bayesian estimator

Modification of notation: (X, Θ) gaussian in $G \times H$, τ prior distribution for Θ , then the Bayesian estimator of θ is

$$E(\Theta | X) = \sum_{i,j} \frac{E(\langle X, v_i \rangle_G \langle \Theta, w_j \rangle_H)}{E(\langle X, v_i \rangle_G^2)} \langle X, v_i \rangle_G w_j$$

- Existence of density not required,
- G (resp. H) may be finite or infinite dimensional.

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Tensorial product

Assume that (X, Y) is gaussian and such that

$$E(Y|X) = l_0(X)$$

where $l_0 \in \mathcal{L}(G, H)$. Thus

$$Y = l_0(X) + \eta$$

where η is strongly orthogonal to X . Then the tensorial product $Y \otimes Y$ has conditional expectation

$$E(Y \otimes Y | X) = l_0(X) \otimes l_0(X) + C_\eta,$$

with

$$l_0(X) = \sum_{i \in I} \frac{\langle X, v_i \rangle_G}{\alpha_i} C_{X, Y}(v_i).$$

Consider a sample (X_i, Y_i) , $1 \leq i \leq n$ and suppose that X_{n+1} is observed. In order to “estimate” $\lambda_0(X_{n+1})$ the following steps are necessary

- Compute the empirical eigenvectors and eigenvalues from

$$C_{n,X} = \frac{1}{n} \sum_{i=1}^n X_i \otimes X_i$$

and

$$C_{n,Y} = \frac{1}{n} \sum_{i=1}^n Y_i \otimes Y_i$$

- Choose a double truncation index
- Find a doctoral student for the calculations.

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